A New and Improved Population-Based Canadian Reference for Birth Weight for Gestational Age

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ABSTRACT. *Background.* **Existing fetal growth references all suffer from 1 or more major methodologic problems, including errors in reported gestational age, biologically implausible birth weight for gestational age, insufficient sample sizes at low gestational age, single-hospital or other non-population–based samples, and inadequate statistical modeling techniques.**

Methods. **We used the newly developed Canadian national linked file of singleton births and infant deaths for births between 1994 and 1996, for which gestational age is largely based on early ultrasound estimates. Assuming a normal distribution for birth weight at each gestational age, we used the expectation-maximization algorithm to exclude infants with gestational ages that were more consistent with 40-week births than with the observed gestational age. Distributions of birth weight at the corrected gestational ages were then statistically smoothed.**

Results. **The resulting male and female curves provide smooth and biologically plausible means, standard deviations, and percentile cutoffs for defining small- and large-for-gestational-age births. Large-for-gestational age cutoffs (90th percentile) at low gestational ages are considerably lower than those of existing references, whereas small-for-gestational-age cutoffs (10th percentile) postterm are higher. For example, compared with the current World Health Organization reference from California (Williams et al, 1982) and a recently proposed US national reference (Alexander et al, 1996), the 90th percentiles for singleton males at 30 weeks are 1837 versus 2159 and 2710 g. The corresponding 10th percentiles at 42 weeks are 3233 versus 3086 and 2998 g.**

Conclusions. **This new sex-specific, population-based reference should improve clinical assessment of growth in individual newborns, population-based surveillance of geographic and temporal trends in birth weight for gestational age, and evaluation of clinical or public health interventions to enhance fetal growth.** *Pediatrics* **2001;108(2). URL: http://www.pediatrics.org/cgi/content/full/108/2/e35;** *fetal growth, birth weight, gestational age, preterm birth, postterm birth.*

ABBREVIATIONS. SD, standard deviation; LGA, large for gestational age; SGA, small for gestational age.

For more than half a century, clinicians and investigators have proposed reference data for assessing birth weight for gestational age. These references have been used by clinicians and researchers to assess fetal growth in individual infants and in populations.[1] Unfortunately, none of the available references are entirely satisfactory. Some are hospital-based,[2–7] giving rise to potential selection bias and problems of generalizability, particularly in view of the low[2,6,7] or high[4,5] socioeconomic status or high altitude[2] that characterizes some of the study hospitals; others are prescriptive rather than descriptive, ie, they are based on infants without known risk factors for impaired fetal growth and thus may not be applicable to populations with mixtures of low- and high-risk pregnancies.[8,9] Some are unisex references that fail to account for the known larger birth weight for gestational age in male versus female infants[3,4,6,10]; others[11,12] go to the opposite extreme and provide curves that are specific for different races, parity, maternal size, and other customizing features for which available data do not permit confident inferences as to whether variations in fetal growth are physiologic or pathologic. Finally, some references are now several decades old and may no longer be pertinent to infants born in more recent years, given the increase in the size of infants born at or near term over the last several decades.

The largest problem with existing references for birth weight for gestational age, however, relates to the measurement of gestational age. Early references rounded gestational age to the nearest week,[2–6] rather than truncated to completed weeks. Although this practice makes sense from a biological standpoint, it is not consistent with World Health Organization (WHO; International Classification of Diseases) recommendations to base gestational age on the number of completed weeks, and references based on the nearest week cannot be applied to populations complying with the WHO recommendations. But even when measured in completed weeks, gestational ages in older references are flawed by

being based on the date of onset of the last menstrual period,[2–6,10,13–16] which has been shown to grossly overestimate the proportion of infants with postterm gestational ages and underestimate those born preterm when compared with early ultrasound measurements.[17] This leads to artifactually wide, or even bimodal, distributions of birth weight at very low gestational ages, and flattening or declining curves postterm, despite the use of various statistical or clinical methods for excluding biologically implausible combinations of birth weight and gestational age.[10,14,16,18,19] Finally, sample sizes for many references are too small, particularly at low gestational ages, leading to irregularity and even nonmonotonicity in the shape of fetal growth curves at these low gestational ages. Some recent references have not smoothed the curves to correct these irregularities.[20,21]

In this study, we have attempted to overcome these previous deficits by constructing a sex-specific reference of birth weight for gestational age that has several advantages over existing references. It is based on a linked live birth-infant death file for all infants born in all of the Canadian provinces and territories (with the exception of Ontario) born between 1994 and 1996, during which time gestational ages have reflected the increasing use of early ultrasound for gestational age estimation. Gestational age is measured in completed weeks, and we have used a new mixture distribution method to exclude infants with implausible gestational ages. This results in smooth, monotonic curves with biologically sensible distributions at all gestational ages. We present both a graphical depiction and a tabular representation of means, standard deviations (SDs), and the 3rd, 5th, 10th, 50th (median), 90th, 95th, and 97th percentiles in the hope that they will be useful to clinicians in classifying fetal growth of newborn infants under their care and of stillbirths, and to researchers and public policy makers in comparing geographic differences and temporal trends in birth weight for gestational age in populations.

METHODS

The references we constructed were based on all births contained in the linked file of live births and infant deaths occurring in the provinces and territories of Canada (with the exception of Ontario) born between January 1, 1994, and December 31, 1996. (The linked file was used, rather than the entire Canadian Birth Database, because the linkage procedure allowed elimination of several duplicates contained in the Birth Database.) Ontario was excluded from the study population base for constructing these references because of documented problems with data quality.[22] Canadian birth certificates include gestational age recorded in completed weeks. Although no information is contained on the birth certificate about the method used to assess gestational age, nor documentation of the dates and results of ultrasound estimates, careful evaluation of these data over several decades strongly suggests that early ultrasound has increasingly been the basis for gestational age assessments in recent years.[23,24] No information on racial origin is contained on Canadian birth certificates.

Our proposed reference is based on singletons with recorded gestational ages of 22 to 43 weeks. The reference sample comprises 347 570 males and 329 035 females. We used a mixture distribution method for correcting gestational age errors, a modification of the procedure described by Oja et al.[25] The model of Oja et al assumes a log normal distribution of birth weight at each gestational age

and errors of +4 weeks or −4 weeks. As assumed in previous studies[12,16] and as confirmed by ultrasound-based estimated fetal weights,[26] our method assumes that the true distribution of birth weight at each gestational age is normal (Gaussian). We also assume that the observed gestational ages are a mixture of correct values and of term (40-week) gestations erroneously measured at the observed (recorded) gestational age. This latter assumption is consistent with the frequently observed curve "bump" at the upper percentiles for extremely preterm infants.[10,14,16]

We used the estimation-maximization algorithm[27] to derive maximum likelihood estimates of the mean and SD and of the probability that the true gestational age of an individual infant was the observed (recorded) gestational age rather than 40 weeks. Starting values for the mean and SD parameters were estimated from the lower half of the observed distribution of birth weight for gestational age (which should not be affected by misclassifying infants born at term as extremely preterm). We then resampled all observations. This resampling resulted in either maintaining the infant in the dataset with the recorded gestational age or deleting that infant, with the probability of deletion equal to the estimated probability that the true gestational age was 40 weeks. Finally, the percentiles of the birth weight distribution at the corrected gestational ages were generated using a smoothing spline with 7 degrees of freedom,[28] weighted by the square root of the (corrected) number of infants at each gestational age.

Tables and graphs were created separately for males and females for the 3rd, 5th, 10th, 50th (median), 90th, 95th, and 97th percentiles at 22 to 43 completed weeks, based on the smoothed estimated curves, and for the mean and SD calculated from the empirical distribution of birth weights after correction. The means and SDs are also tabulated to allow calculation of z scores rather than (or in addition to) percentiles, where

$$z = \frac{\text{observed birth weight} - \text{mean birth weight}}{\text{SD}}$$

RESULTS

Tables 1 and 2 show the 3rd, 5th, 10th, 50th (median), 90th, 95th, and 97th percentile; mean; and SD birth weights for male and female infants, respectively. Figures 1 and 2 show the crude percentile graphs, and Figs 3 and 4 show the curves produced after correcting the gestational ages and smoothing the resulting birth weight distributions. The crude curves show a convex "bump" in the upper percentiles at 27 to 33 weeks, consistent with our model's assumption that some truly term births are misclassified at these (grossly underestimated) gestational ages. The corrected curves are smooth and evenly spaced and show no biologically implausible bumps preterm or flattening or declines postterm. Moreover, each of the percentile curves shows the expected sigmoid shape, with birth weight increasing monotonically as gestational age advances. The smoothed curves closely follow (within limits of random error) the empiric birth weight distribution at each (corrected) gestational age. For example, at 30 weeks, the 90th percentile curve cuts off 91.4% of males and 87.7% of females in our data set. At 42 weeks, the 10th percentile curve cuts off 10.4% of males and 10.2% of females.

Figure 5 illustrates how our procedure has corrected the crude curves at 30 weeks' gestation (for males). At low percentiles, the curves before and after correction are virtually identical, but starting at the 90th percentile, and especially at the 95th and 97th percentiles, the cutoffs after the correction are substantially lower.

Table 3 compares our new proposed reference (based on the corrected curves) with 4 population-

TABLE 1. Birth Weight (g) for Gestational Age, Canadian Male Singletons Born Between 1994 and 1996, Corrected and Smoothed

Gestational Age	n*	3rd Percentile	5th Percentile	10th Percentile	50th Percentile	90th Percentile	95th Percentile	97th Percentile	Mean	SD
22	82	338	368	401	490	587	627	659	501	111
23	114	406	434	475	589	714	762	797	598	114
24	156	468	498	547	690	844	902	940	697	125
25	202	521	557	617	795	981	1048	1092	800	147
26	234	571	614	686	908	1125	1200	1251	909	178
27	254	627	677	763	1033	1278	1358	1416	1026	209
28	330	694	752	853	1173	1445	1532	1598	1159	241
29	392	780	845	964	1332	1629	1729	1809	1312	273
30	467	885	959	1099	1507	1837	1955	2053	1487	306
31	584	1012	1098	1259	1698	2069	2209	2327	1682	339
32	997	1164	1266	1444	1906	2319	2478	2614	1896	369
33	1368	1344	1460	1648	2127	2580	2750	2897	2123	391
34	2553	1552	1677	1866	2360	2851	3029	3184	2361	410
35	4314	1783	1907	2091	2600	3132	3318	3475	2607	428
36	9648	2024	2144	2321	2845	3411	3604	3759	2855	443
37	19 965	2270	2384	2552	3080	3665	3857	4003	3091	449
38	51 947	2498	2605	2766	3290	3877	4065	4202	3306	448
39	77 623	2684	2786	2942	3465	4049	4232	4361	3489	445
40	112 737	2829	2927	3079	3613	4200	4382	4501	3638	447
41	54 139	2926	3025	3179	3733	4328	4512	4631	3745	459
42	8791	2960	3070	3233	3815	4433	4631	4773	3800	485
43	276	2954	3081	3249	3864	4528	4747	4941	3793	527

* Sample size at each gestational age after exclusions.

TABLE 2. Birth Weight (g) for Gestational Age, Canadian Female Singletons Born Between 1994 and 1996, Corrected and Smoothed

Gestational Age	n*	3rd Percentile	5th Percentile	10th Percentile	50th Percentile	90th Percentile	95th Percentile	97th Percentile	Mean	SD
22	80	332	347	385	466	552	576	576	472	72
23	106	379	403	450	557	669	706	726	564	95
24	148	424	456	513	651	790	839	887	656	121
25	184	469	508	578	751	918	982	1060	754	152
26	191	516	562	645	858	1060	1139	1247	860	186
27	188	569	624	717	976	1218	1313	1446	976	222
28	287	634	697	802	1109	1390	1499	1657	1107	254
29	299	716	787	903	1259	1578	1701	1885	1256	286
30	390	814	894	1022	1427	1783	1918	2121	1422	319
31	461	938	1026	1168	1613	2004	2150	2347	1604	345
32	795	1089	1184	1346	1817	2242	2399	2578	1808	368
33	1055	1264	1369	1548	2035	2494	2664	2825	2029	389
34	2018	1467	1581	1768	2266	2761	2948	3097	2266	409
35	3391	1695	1813	1998	2506	3037	3242	3384	2512	426
36	8203	1935	2052	2227	2744	3307	3523	3660	2754	439
37	17 308	2177	2286	2452	2968	3543	3752	3886	2981	443
38	47 516	2406	2502	2658	3169	3738	3931	4061	3181	439
39	75 068	2589	2680	2825	3334	3895	4076	4202	3350	434
40	110 738	2722	2814	2955	3470	4034	4212	4331	3486	434
41	52 063	2809	2906	3051	3576	4154	4330	4444	3588	439
42	7970	2849	2954	3114	3655	4251	4423	4554	3656	448
43	277	2862	2975	3159	3717	4333	4495	4685	3693	459

* Sample size at each gestational age after exclusions.

based references: 1 from Canada (Arbuckle et al[20]) and 3 from the United States (Williams et al[16] [white non-Hispanics only], Zhang and Bowes[12] [whites only], and Alexander et al[19]). Tabulated values for the Canadian reference were provided by Statistics Canada (Russell Wilkins, Health Analysis and Modeling Group, Statistics Canada, Ottawa, Ontario, personal communication). The comparison focuses on the 2 sensitive regions mentioned earlier: the 90th percentile (the conventional large-for-gestational-age [LGA] cutoff) at 30 weeks, and the 10th percentile (the conventional small-for-gestational-age [SGA] cutoff) at 42 weeks. Each of the previous references uses a correction procedure to reduce errors in pre-term gestational age estimates, but each is characterized by residual bumps preterm and/or flattening or declining curves postterm. Our curves yield the lowest LGA cutoff at 30 weeks, with differences as large as 900 g. Our curves also yield the highest SGA cutoff at 42 weeks, with differences up to 400 g. All 4 of our cutoffs, however, are close to those of the earlier Canadian reference by Arbuckle et al.[20] The mean values also differ at these gestational ages. Our reference mean for males is 1487 g at 30 weeks and 3800 g at 42 weeks. The corresponding figures for Williams et al[16] are 1537 and 3665 g; for Zhang and Bowes,[12] 1653 and 3548 g; and for Alexander et al,[19] 1637 and 3522 g.

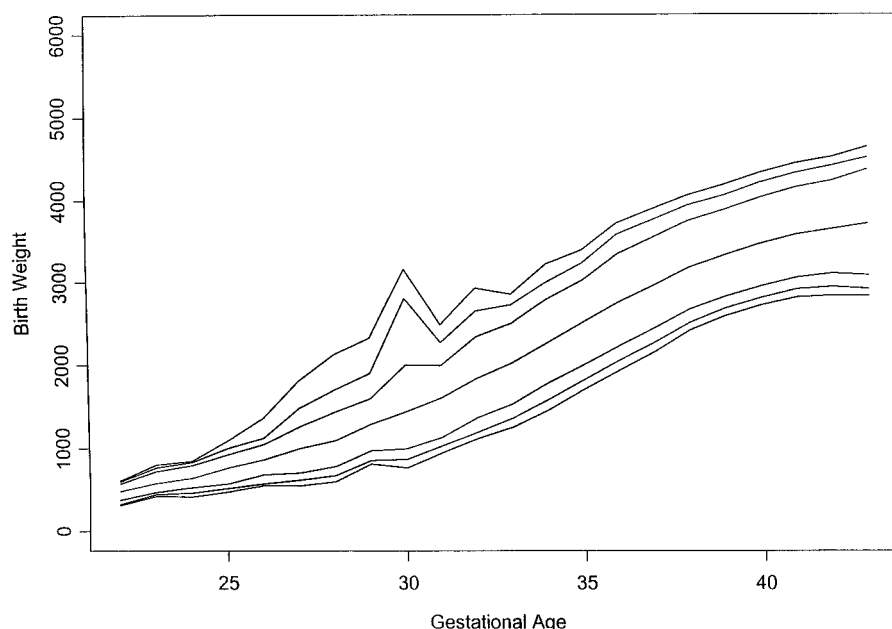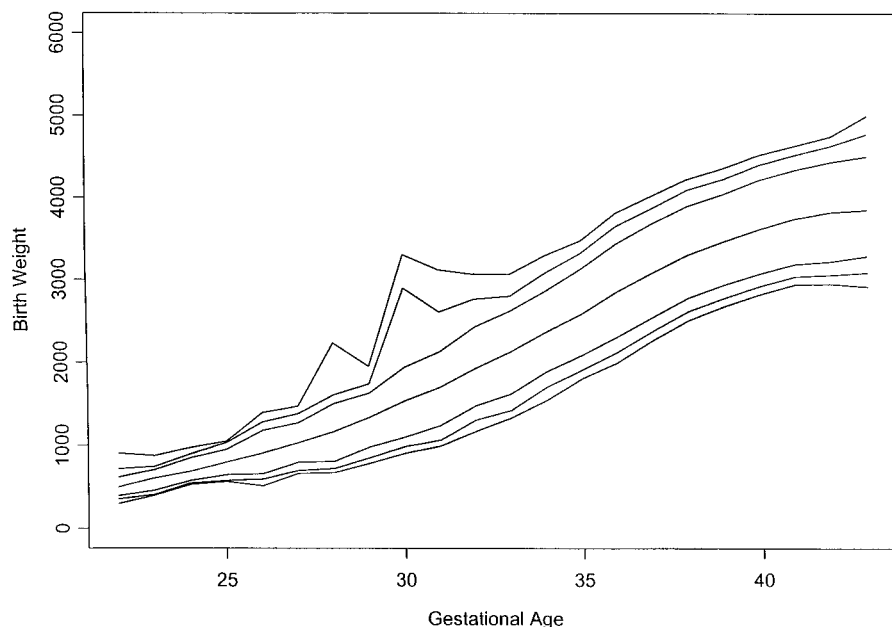**Fig 1.** Canadian male singletons, crude curves.

Fig 2. Canadian female singletons, crude curves.

DISCUSSION

The reference curves and tables presented herein differ in several important respects from those currently available. Their main advantages include their recent vintage; population base; increased reliance on early ultrasound-based estimates of gestational age; statistical modeling of gestational ages to correct biologically implausible values (based on the observed and expected birth weights); sex-specificity; and adaptability to the use of either percentiles or *z* scores. This reference differs from others currently available in showing lower means and particularly upper (90th, 95th, and 97th) percentile birth weights at low gestational ages, as well as higher birth weights (particularly at the 10th, 5th, and 3rd percentiles) for gestational age in the postterm period.

The absence of a downturn in the curves in the postterm period is consistent with evidence based on early ultrasound-based gestational ages.[29,30] Our reference also differs from older references by reflecting the recent temporal trend toward increasing birth weight at or near term.

This new reference should prove helpful for several types of users in developed country settings. For clinicians, it enables better classification of individual infants as small, appropriate, or large for their gestational age. It should prove useful for clinicians caring for extremely preterm infants in whom unexpectedly high birth weights should lead to skepticism about the reported gestational age. It should also be reassuring to prenatal care providers by eliminating the impression that fetuses fail to gain, or

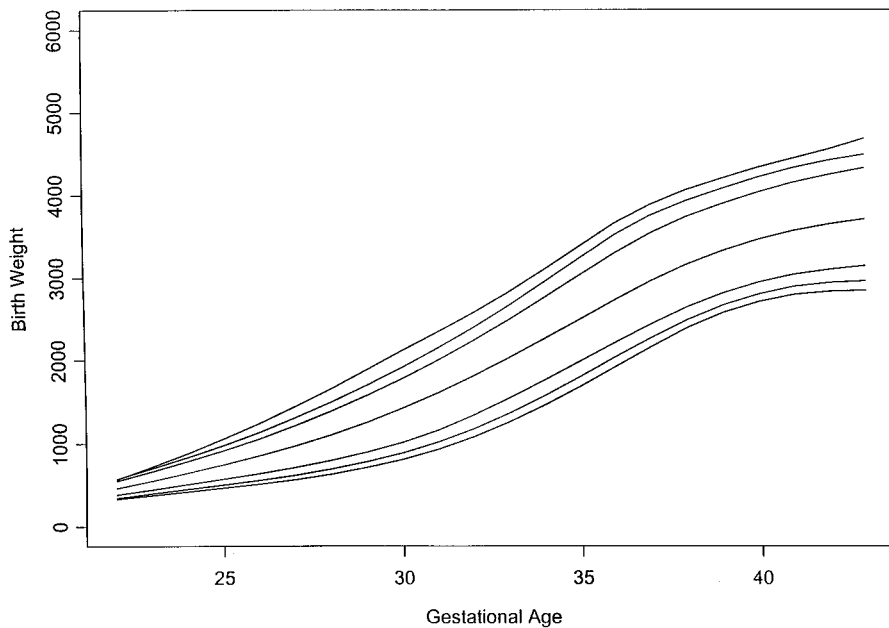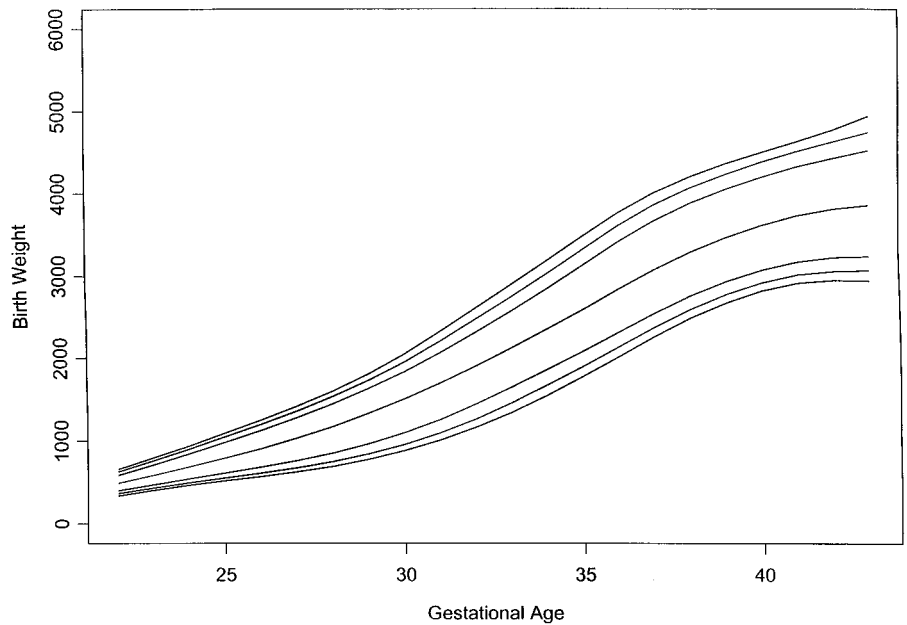**Fig 3.** Canadian male singletons, corrected and smoothed curves.

Fig 4. Canadian female singletons, corrected and smoothed curves.

even lose, weight postterm. For researchers, this new reference should allow better classification of groups of infants for determination of geographic differences, temporal trends, etiologic determinants, and short- and long-term prognosis. For public health policy makers, the reference should be useful in surveillance, ie, in tracking population differences by geographic location, socioeconomic status, race/ethnicity, and other relevant factors, as well as in tracking trends over time and in response to educational or other public health interventions.

Despite the advantages of our new reference over its predecessors, several limitations should be acknowledged. First, like any purely statistical correction procedure, our method compares poorly with early ultrasound for valid estimation of the true gestational age.[17] In particular, early ultrasound corrects

last menstrual period-based estimates by a systematic shift to the left (ie, toward lower gestational ages) across the entire gestational age distribution, primarily because of the effect of delayed ovulation.[31] Our method, however, corrects only gross errors by excluding preterm (and some postterm) infants whose birth weights are more compatible with 40-week births. We have also experimented with more complex statistical corrections (eg, ±1, 2, 4, and 8 weeks), but the more complex algorithms are far more computer-intensive, result in many more exclusions, and do not substantially alter the slopes or SGA or LGA cutoffs of the reference curves.

The most important limitation of all population-based references, including ours, is their cross-sectional nature, ie, they are based on the birth weights of different infants born at different gestational ages,

Fig 5. Birth weight distribution at 30 weeks, male singletons, before and after correction.
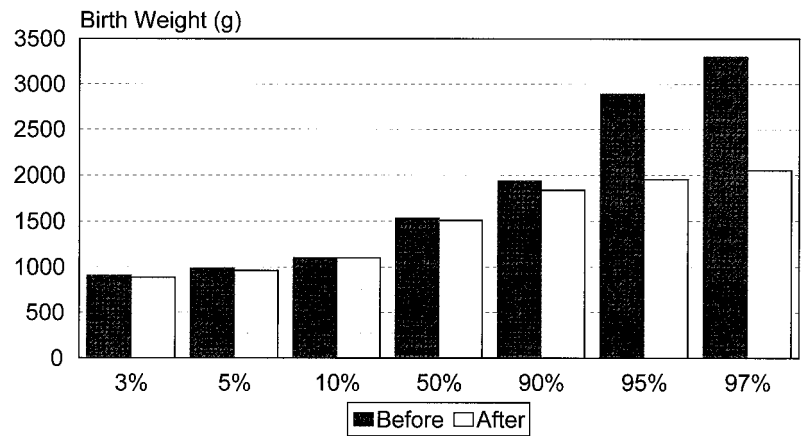
TABLE 3. Comparison of Key Birth Weight Cutoffs (g), Proposed New Reference Versus Four Previous References

Reference	LGA (90th Percentile) at 30 Weeks		SGA (10th Percentile) at 42 Weeks	
	Male	Female	Male	Female
Proposed reference	1837	1783	3233	3114
Arbuckle et al[20]	1870	1800	3200	3070
Williams et al[16]	2159	2113	3086	2936
Zhang and Bowes[12]	2355*	2288*	2835*	2735*
Alexander et al[19]	2710†	2710†	2998	2884

* Average of values for primiparas and multiparas.
† Sex-specific value not reported.

rather than longitudinal measurements of the same infants over the course of gestation.[32] Growth is defined as an increase in size over time, and documentation of increasing size therefore requires 2 or more serial measurements. Unfortunately, serial anthropometric measurements during fetal life are feasible only with ultrasound, and these have not proved to be sufficiently valid or precise to serve as a reference.[33–35] The substitution of cross-sectional for longitudinal data on fetal growth is a problem, because evidence suggests that preterm infants are somewhat smaller than fetuses of the same gestational age who remain in utero.[36–39] This problem may partly reflect the fact that some of the determinants of fetal growth and gestational duration overlap. Pregnancy-induced hypertension and preeclampsia, for example, are risk factors for both intrauterine growth restriction and both spontaneous and induced preterm delivery.[39–42] At or after term, fetuses who remain unborn may not have grown at the same rate as those born earlier; because fetal size is considered to be one of the determinants of the onset of labor, cross-sectional data may reflect the earlier birth of faster-growing fetuses.[1]

If and when ultrasound technology improves to permit more valid and reliable measurements of fetal weight and other anthropometric measurements in utero, a truly longitudinal fetal growth reference will be feasible. Other future changes may also require modification or replacement of our reference. Temporal trends toward increasing maternal prepregnancy weight and weight gain during pregnancy, maternal stature, and reductions in cigarette smok-

ing will probably continue to increase the size of infants born at or near term. Finally, the increasing trend toward obstetric intervention to hasten delivery for pathologic pregnancies in the preterm period, and in response to signs of slow growth or other fetal compromise in the postterm period, may continue to affect the shape of the growth curve at these periods of gestation.

REFERENCES

1. World Health Organization Expert Committee on the Use and Interpretation of Anthropometry. *Physical Status: The Use and Interpretation of Anthropometry*. Geneva, Switzerland: World Health Organization; 1995
2. Lubchenco L, Hansman C, Dressler M, Boyd E. Intrauterine growth as estimated from liveborn birth weight data at 24 to 42 weeks of gestation. *Pediatrics*. 1963;32:793–800
3. Gruenwald P. Growth of the human fetus. I. Normal growth and its variation. *Am J Obstet Gynecol*. 1966;94:1112–1119
4. Usher R, McLean F. Intrauterine growth of live-born Caucasian infants at sea-level: standards obtained from measurements in 7 dimensions of infants born between 25 and 44 weeks gestation. *J Pediatr*. 1969;74:901–910
5. Babson S, Behrman R, Lessel R. Liveborn birth weights for gestational age of white middle class infants. *Pediatrics*. 1970;45:937–943
6. Brenner W, Edelman D, Hendricks C. A standard of fetal growth for the United States of America. *Am J Obstet Gynecol*. 1976;126:555–564
7. Amini S, Catalano P, Hirsch V, Mann L. An analysis of birth weight by gestational age using a computerized perinatal data base, 1975–1992. *Obstet Gynecol*. 1994;83:342–352
8. Lawrence C, Fryer J, Karlberg J, Niklasson A, Ericson A. Modelling of reference values for size at birth. *Acta Paediatr Scand*. 1989;350(suppl):55–69
9. Niklasson A, Ericson A, Fryer J, Karlberg J, Lawrence C, Karlberg P. An update of the Swedish reference standards for weight, length, and head circumference at birth for given gestational age (1977–1981). *Acta Paediatr Scand*. 1991;89:756–762
10. David R. Population-based intrauterine growth curves from computerized birth certificates. *South Med J*. 1983;76:1401–1406
11. Gardosi J, Chang A, Kalyan B, Sahota D, Symonds E. Customized antenatal growth charts. *Lancet*. 1992;339:283–287
12. Zhang J, Bowes W Jr. Birth-weight-for-gestational-age patterns by race, sex, and parity in the United States population. *Obstet Gynecol*. 1995;86:200–208
13. Gibson J, McKeown T. Observations of all births (23,970) in Birmingham, 1947;VI:Birth weight, duration of gestation, and survival related to sex. *Br J Soc Med*. 1952;6:152–158
14. Battaglia F, Frazier T, Hellegers A. Birth weight, gestational age, and pregnancy outcome, with special reference to high birth weight-low gestational age infant. *Pediatrics*. 1966;37:417–422
15. Thomson A, Billewicz W, Hytten F. The assessment of fetal growth. *J Obstet Gynaecol Br Common*. 1968;75:903–916
16. Williams R, Creasy R, Cunningham G, Hawes W, Norris F, Tashiro M. Fetal growth and perinatal viability in California. *Obstet Gynecol*. 1982;59:624–632
17. Kramer M, McLean F, Boyd M, Usher R. The validity of gestational age

estimation by menstrual dating in term, preterm, postterm gestations. *JAMA*. 1988;260:3306–3308

18. Källén B. A birth weight for gestational age standard based on data in the Swedish Medical Birth Registry, 1985–1989. *Eur J Epidemiol*. 1995;11: 601–606

19. Alexander G, Himes J, Kaufman R, Mor J, Kogan M. A United States national reference for fetal growth. *Obstet Gynecol*. 1996;87:163–168

20. Arbuckle T, Wilkins R, Sherman G. Birth weight percentiles by gestational age in Canada. *Obstet Gynecol*. 1993;81:39–48

21. Roberts C, Mueller L, Hadler J. Birth weight percentiles by gestational age, Connecticut 1988–1993. *Conn Med*. 1996;60:131–140

22. Joseph K, Kramer M. Recent trends in infant mortality rates and proportions of low-birth-weight live births in Canada. *Can Med Assoc J*. 1997;157:535–541

23. Arbuckle T, Sherman G. An analysis of birth weight by gestational age in Canada. *Can Med Assoc J*. 1989;140:157–165

24. Joseph K, Kramer M. Recent versus historical trends in preterm birth in Canada. *Can Med Assoc J*. 1999;161:1409

25. Oja H, Koiranen M, Rantakallio P. Fitting mixture models to birth weight data: a case study. *Biometrics*. 1991;47:883–897

26. Marsál K, Persson P-H, Larsen T, Lilja H, Selbing A, Sultan B. Intrauterine growth curves based on ultrasonically estimated foetal weights. *Acta Paediatr*. 1996;85:843–848

27. Dempster A, Laird N, Rubin D. Maximum likelihood from incomplete data via the EM algorithm. *J R Stat Soc, Series B*. 1997;39:1–38

28. Hastie T, Tibshirani R. *Generalized Additive Models*. New York, NY: Chapman & Hall/CRC; 1990

29. McLean F, Boyd M, Usher R, Kramer M. Postterm infants: too big or too small? *Am J Obstet Gynecol*. 1991;164:619–624

30. Wilcox M, Gardosi J, Mongelli M, Ray C, Johnson I. Birth weight from pregnancies dated by ultrasonography in a multicultural British population. *Br Med J*. 1993;307:588–591

31. Yang H, Kramer M, Platt R, et al. How does early ultrasound estimation of gestational age lead to higher rates of preterm birth? *Am J Obstet Gynecol*. 2001; in press

32. Wilcox A. Birth weight, gestation, and the fetal growth curve. *Am J Obstet Gynecol*. 1981;139:863–867

33. Little G. Fetal growth and development. In: Eden R, Boehm F, eds. *Assessment and Care of the Fetus: Physiological, Clinical, and Medico-Legal Principles*. Norwalk, CT: Appleton & Lange; 1990:1–5

34. Rogers M, Chung T, Chang A. Ultrasound fetal weight estimation: precision or guess work? *Aust N Z J Obstet Gynaecol*. 1993;33:142–144

35. Chauhan S, Charania S, McLaren R, et al. Ultrasonographic estimate of birth weight at 24 to 34 weeks: a multicenter study. *Am J Obstet Gynecol*. 1998;179:916

36. Ott W, Doyle S. Normal ultrasonic fetal weight curve. *Obstet Gynecol*. 1982;59:603–606

37. Weiner C, Sabbagha R, Vaisrub N, Depp R. A hypothetical model suggesting suboptimal intrauterine growth in infants delivered preterm. *Obstet Gynecol*. 1985;65:323–326

38. Secher N, Hansen P, Thomsen B, Keiding N. Growth retardation in preterm infants. *Br J Obstet Gynaecol*. 1987;94:115–120

39. Hediger M, Scholl T, Schall J, Miller L, Fischer R. Fetal growth and the etiology of preterm delivery. *Obstet Gynecol*. 1995;85:175–182

40. Kramer M. Determinants of low birth weight: methodological assessment and meta-analysis. *Bull World Health Organ*. 1987;65:663–737

41. Ananth C, Savitz D, Luther E, Bowes W Jr. Preeclampsia and preterm birth subtypes in Nova Scotia, 1986 to 1992. *Am J Perinatol*. 1997;14:17–23

42. Kramer M, Séguin L, Lydon J, Goulet L. Socio-economic disparities in pregnancy outcome: why do the poor fare so poorly? *Paediatr Perinat Epidemiol*. 2000;14:194–210

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Michael S. Kramer, Robert W. Platt, Shi Wu Wen, K.S. Joseph, Alexander Allen,
Michal Abrahamowicz, Béatrice Blondel, Gérard Bréart and for the Fetal/Infant
Health Study Group of the Canadian Perinatal Surveillance System

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